# Deep Generative Models 

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DeepMind

## 2-minute exercise

Talk to your friend next to you, and tell him or her everything you can about this data set:


## Data



## Data manifold

We can capture most of the variability in the data through one number

$$
z^{(n)}=1 \text { or } 2,3,4
$$

for each image $n$, even though each image is 16 dimensional

## How?

## How?

## 1. $\quad$ Take $z^{(n)}=2$

2. Draw bar in column 2 of image
3. Et voila! You have $x^{(n)}$
$z^{(n)}=2$

$x^{(n)}$


## How?

## 1. Take $z^{(n)}=2$

2. Draw bar in column 2 of image
3. Et voila! You have $x^{(n)}$
$z^{(n)}=2$

## Maybe some neural network, that takes $z$ as input, and outputs a 16-dimensional vector x...?

$x^{(n)}$

## 3-minute exercise

Write or draw a function (like a multi-layer perceptron) that takes $z \in \mathbb{R}$ and produces $\boldsymbol{X}$

Is your input one-dimensional?
Is your output 16-dimensional?
Identify all the "tunable" parameters $\boldsymbol{\theta}$ of your function

$$
z^{(n)}=2
$$



$$
x^{(n)}
$$



## 3-minute exercise

Write or draw a function (like a multi-layer perceptron) that takes $z \in \mathbb{R}$ and produces $\boldsymbol{X}$

Is your input one-dimensional?
Is your output 16-dimensional?
Identify all the "tunable" parameters $\boldsymbol{\theta}$ of your function

## Data manifold

The 16-dimensional images live on a 1-dimensional manifold, plus some "noise"


## ...and noise

The 16-dimensional images live on a 1 -dimensional manifold, plus some "noise"


## 3-minute exercise

Change your multi-layer perceptron to take $\boldsymbol{Z}$ and produce a distribution over $\boldsymbol{X}$

$$
p_{\theta}(x \mid z)
$$



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Change your multi-layer perceptron to take $\boldsymbol{Z}$ and produce a distribution over $\boldsymbol{X}$

$$
p_{\theta}(x \mid z)
$$

## Decoder

```
def generative_network(z, ...):
```

.
return bernoulli_logits \# for binary pixels

$z \rightarrow x$

Inference

$?$

## Inversing our world

Two BIG problems to solve:
Inference
You wrote down $P_{\theta}(x \mid z)$ and can compute it.
Say I give you $\boldsymbol{X}$. Keeping $\theta$ fixed, what was $\mathcal{Z}$ ? Or $p_{\theta}(z \mid x)$ ?

## Learning

Is there a better (best) $\theta$ to generate the observed $\boldsymbol{X}$ from $\boldsymbol{Z}$ ?

## Inference

You wrote down $p_{\theta}(x \mid z)$ and can compute it.
Say I give you $\boldsymbol{X}$. Keeping $\theta$ fixed, what was $Z$ ? Or $p_{\theta}(z \mid x)$ ?
$z \in \mathbb{R}$ $\qquad$

100001100002100003

## Inference

You wrote down $p_{\theta}(x \mid z)$ and can compute it.
Say I give you $\boldsymbol{X}$. Keeping $\theta$ fixed, what was $Z$ ? Or $p_{\theta}(z \mid x)$ ?
$z \in \mathbb{R}$


100001100002100003

To really answer that question, we need some notion of where we might have started! No inference without prior assumptions :)

## Prior assumptions



## Inference

$$
p(z)=\mathcal{N}(z ; 0,1)
$$

I give you $\boldsymbol{X}$. Keeping $\theta$ fixed, what was $\boldsymbol{Z}$ ?

$$
p_{\theta}(x \mid z)
$$



## Inference

$$
p(z)=\mathcal{N}(z ; 0,1)
$$

I give you $\boldsymbol{X}$. Keeping $\theta$ fixed, what was $\boldsymbol{Z}$ ?

$$
p_{\theta}(x \mid z)
$$

## 3-minute exercise

$$
p(z)=\mathcal{N}(z ; 0,1)
$$

Assuming the largest value of $p_{\theta}(x \mid z)$ is 1 , draw

$$
p_{\theta}(x, z)=p_{\theta}(x \mid z) p(z)
$$

as a function of $\boldsymbol{Z}$ on the same axis as above

Joint density (with x observed)

$$
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Assuming the largest value of $p_{\theta}(x \mid z)$ is 1 , draw

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$$

as a function of $\boldsymbol{Z}$ on the same axis as above

Joint density (with x observed)


Area $=$ ?

1-minute exercise:
what is the area?

## Marginal likelihood (evidence)

## Area $=1$

Area $=$ ?

$$
\begin{aligned}
\text { area } & =\int p_{\theta}(x \mid z) p(z) \mathrm{d} z \\
& =\int p_{\theta}(x, z) \mathrm{d} z \\
& =p_{\theta}(x)
\end{aligned}
$$

## Relationship to posteror



Area $=1$

$$
p_{\theta}(z \mid x)=\frac{p_{\theta}(x \mid z) p(z)}{p_{\theta}(x)} \quad \begin{gathered}
\text { Dividing by the marginal } \\
\text { ilkeihood } \\
\text { area back evidence to } 1 . . .
\end{gathered}
$$

Evidence, for all data points

$$
\begin{gathered}
\left.X \equiv x^{(1)}, x^{(2)} \ldots, x^{(N)} \begin{array}{c}
\text { Area for data } \\
\text { point } n
\end{array}\right) \\
p_{\theta}(X)=\prod_{n=1}^{N} p_{\theta}\left(x^{(n)}\right)
\end{gathered}
$$

Evidence, for all data points

$$
\begin{aligned}
& X \equiv x^{(1)}, x^{(2)} \ldots, x^{(N)} \begin{array}{c}
\text { Area for data } \\
\text { point } n
\end{array} \\
& X=\sum_{n=1}^{N} \log p_{\theta}\left(x^{(n)}\right)
\end{aligned}
$$



## Maximizing the evidence

The product of the areas


These Z's don't generate images like the ones in the data set...
(With this $\boldsymbol{\theta}$, the prior doesn't capture the data manifold well)

## Maximizing the evidence



## For the sharp-sighted



## Learning

We want to maximize

except that we cannot write down an analytically tractable expression for the area.

Strategies: Stochastic (Monte Carlo samples + gradients) or deterministic (approximate inference). We'll follow the "deterministic" path next..

## Approximate inference

We want to use this quantity for "learning", but cannot compute it in an analytically tractable way:

$$
\begin{aligned}
\log p_{\theta}(x) & =\log \int p_{\theta}(x \mid z) p(z) \mathrm{d} z \\
& =\log \int p_{\theta}(x, z) \mathrm{d} z
\end{aligned}
$$

## "Variational lower bound"

Unnormalized $p_{\theta}(z \mid x)$ We cannot (tractably) compute Unnormalized $q_{\phi}(z \mid x)$

Strategy: we choose some other $q_{\phi}(z \mid x)$ so that we can compute the area underneath the blue curve (e.g. Gaussian)


## Encoder

```
def inference_network(x, latent_dim=1):
    return mu, sigma
```


## Encoder decoder



## Strategy

Change $\phi$ to inflate the area under the blue curve. We can do that!
Change $\theta$ to change the green curve, so that we can inflate the area under the blue curve even more
...and so, hopefully, the area under the green curve also gets bigger


## Whhaaaatttt?

## Strategy

## Change $\phi$ to inflate the area under the blue curve. We can do that!

Change $\theta$ to change the green curve, so that we can inflate the area under the blue curve even more
...and so, hopefully, the area under the green curve also gets bigger


## 3-minute exercise

Create and draw $q_{\phi}(z \mid x)=\mathcal{N}\left(z ; \mu_{\phi}(x), \sigma_{\phi}^{2}(x)\right)$ as a function. It could be a multi-layer perceptron (MLP) that takes 16-dimensional $\boldsymbol{X}$, and produces two 1-dimensional quantities,

$$
\begin{aligned}
\text { mean } & =\mu_{\phi}(x) \\
\text { variance } & =\sigma_{\phi}^{2}(x)
\end{aligned}
$$



What are your parameters $\phi$ ?

## Objective function discussion

maximize (for all data points)...


$$
\mathbb{E}_{q_{\phi}(z \mid x)}\left[\log p_{\theta}(x \mid z)\right]-\operatorname{KL}\left(q_{\phi}(z \mid x) \| p(z)\right)
$$

## Evidence lower bound (ELBO) for one data point

$$
\begin{aligned}
\log p_{\theta}(x) & =\log \int p_{\theta}(x \mid z) p(z) \mathrm{d} z \\
& =\log \int q_{\phi}(z \mid x)\left[\frac{p_{\theta}(x \mid z) p(z)}{q_{\phi}(z \mid x)}\right] \mathrm{d} z \\
& \geq \int q_{\phi}(z \mid x) \log \left[\frac{p_{\theta}(x \mid z) p(z)}{q_{\phi}(z \mid x)}\right] \mathrm{d} z \quad \text { JJensen] } \\
& =\int q_{\phi}(z \mid x) \log p_{\theta}(x \mid z) \mathrm{d} z-\int q_{\phi}(z \mid x) \log \left[\frac{q_{\phi}(z \mid x)}{p(z)}\right] \mathrm{d} z \\
& =\mathbb{E}_{q_{\phi}(z \mid x)}\left[\log p_{\theta}(x \mid z)\right]-\mathrm{KL}\left(q_{\phi}(z \mid x) \| p(z)\right) \\
\mathrm{ELBO} \longrightarrow & \equiv \mathcal{L}(x ; \theta, \phi)
\end{aligned}
$$

## Evidence lower bound (ELBO) for one data point

$$
\log p_{\theta}(x)=\log \int p_{\theta}(x \mid z) p(z) \mathrm{d} z
$$

$$
=\log \int q_{\phi}(z \mid x)\left[\frac{p_{\theta}(x \mid z) p(z)}{q_{\phi}(z \mid x)}\right] \mathrm{d} z
$$

## Evidence lower bound (ELBO) for one data point

$$
\log p_{\theta}(x)=\log \int p_{\theta}(x \mid z) p(z) \mathrm{d} z
$$

concave function

$$
\begin{aligned}
& =\log \int q_{\phi}(z \mid x)\left[\frac{p_{\theta}(x \mid z) p(z)}{q_{\phi}(z \mid x)}\right] \underline{\mathrm{d} z} \\
& \geq \int q_{\phi}(z \mid x) \log \left[\frac{p_{\theta}(x \mid z) p(z)}{q_{\phi}(z \mid x)}\right] \frac{\mathrm{d} z}{\text { [Jensen] }} \\
& \text { expectation concave function }
\end{aligned}
$$

$$
\log \mathbb{E}_{q_{\phi}(z \mid x)}[f(z)] \geq \mathbb{E}_{q_{\phi}(z \mid x)}[\log f(z)]
$$

## 3-minute exercise

Jensen's inequality
Draw $\log (\ldots)$ as a function, and convince yourself that

$$
\log \left(\underline{\frac{2}{3}} z_{1}+\frac{1}{3} z_{2}\right) \geq \underline{\frac{2}{3}} \log \left(z_{1}\right)+\frac{1}{3} \log \left(z_{2}\right)
$$

is always true for any (nonnegative) setting of $z_{1}$ and $z_{2}$.

## Logarithm (concave)

$$
\log \left(\frac{2}{3} z_{1}+\frac{1}{3} z_{2}\right) \geq \frac{2}{3} \log \left(z_{1}\right)+\frac{1}{3} \log \left(z_{2}\right)
$$



## Evidence lower bound (ELBO) for one data point

$$
\log p_{\theta}(x)=\log \int p_{\theta}(x \mid z) p(z) \mathrm{d} z
$$

| Reconstruction |
| :---: |
| Expected log likelihood. |
| Cannot compute in |
| closed form, and will |
| have to get a Monte |
| Carlo estimate |
| (with SGD) |

$$
=\log \int q_{\phi}(z \mid x)\left[\frac{p_{\theta}(x \mid z) p(z)}{q_{\phi}(z \mid x)}\right] \mathrm{d} z
$$

$$
\begin{aligned}
& \geq \int q_{\phi}(z \mid x) \log \left[\frac{p_{\theta}(x \mid z) p(z)}{q_{\phi}(z \mid x)}\right] \mathrm{d} z \quad \begin{array}{c}
\text { distributions (here). } \\
\text { Can compute in closed } \\
\text { form }
\end{array} \\
& =\int q_{\phi}(z \mid x) \log p_{\theta}(x \mid z) \mathrm{d} z-\int q_{\phi}(z \mid x) \log \left[\frac{q_{\phi}(z \mid x)}{p(z)}\right] \mathrm{d} z \\
& =\mathbb{E}_{q_{\phi}(z \mid x)}\left[\log p_{\theta}(x \mid z)\right]-\operatorname{KL}\left(q_{\phi}(z \mid x) \| p(z)\right) \\
& \equiv \mathcal{L}(x ; \theta, \phi)
\end{aligned}
$$

Kullback-Leibler divergence between two Gaussian

## Expected log likelihood

We can estimate the expected log likelihood with a Monte Carlo estimate:
Draw L samples $z^{(l)} \sim \mathcal{N}\left(z ; \mu_{\phi}(x), \sigma_{\phi}^{2}(x)\right) \ldots$


## Expected log likelihood

We can estimate the expected log likelihood with a Monte Carlo estimate:
Draw $L$ samples $z^{(l)} \sim \mathcal{N}\left(z ; \mu_{\phi}(x), \sigma_{\phi}^{2}(x)\right)$ and use them to estimate the average:

$$
\begin{aligned}
\mathbb{E}_{q_{\phi}(z \mid x)}\left[\log p_{\theta}(x \mid z)\right] & =\mathbb{E}_{z \sim \mathcal{N}\left(z ; \mu_{\phi}(x), \sigma_{\phi}^{2}(x)\right)}\left[\log p_{\theta}(x \mid z)\right] \\
& \approx \frac{1}{L} \sum_{l=1}^{L} \log p_{\theta}\left(x \mid z^{(l)}\right)
\end{aligned}
$$

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& \approx \frac{1}{L} \sum_{l=1}^{L} \log p_{\theta}\left(x z^{(l)}\right.
\end{aligned}
$$

Using samples in this way removes $\phi$ from part of the objective function, and even though we can evaluate it, we can't take derivatives / get the gradients!

## Naive sampling



## Expected log likelihood: reparameterization trick

We can estimate the expected log likelihood with a Monte Carlo estimate:
Draw L samples $\epsilon^{(l)} \sim \mathcal{N}(\epsilon ; 0,1)$ and transform them!


## Expected log likelihood

We can estimate the expected log likelihood with a Monte Carlo estimate:
Draw $L$ samples $\epsilon^{(l)} \sim \mathcal{N}(\epsilon ; 0,1)$ and use them to estimate the average:

$$
\begin{aligned}
\mathbb{E}_{q_{\phi}(z \mid x)}\left[\log p_{\theta}(x \mid z)\right] & =\mathbb{E}_{\epsilon \sim \mathcal{N}(\epsilon ; 0,1)}\left[\log p_{\theta}\left(x \mid z=\mu_{\phi}(x)+\epsilon * \sigma_{\phi}(x)\right)\right] \\
& \approx \frac{1}{L} \sum_{l=1}^{L} \log p_{\theta}\left(x \mid z^{(l)}=\mu_{\phi}(x)+\epsilon_{\dagger}^{(l)} * \sigma_{\phi}(x)\right)
\end{aligned}
$$

## Expected log likelihood

We can estimate the expected log likelihood with a Monte Carlo estimate:
Draw $L$ samples $\epsilon^{(l)} \sim \mathcal{N}(\epsilon ; 0,1)$ and use them to estimate the average:

$$
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\mathbb{E}_{q_{\phi}(z \mid x)}\left[\log p_{\theta}(x \mid z)\right] & =\mathbb{E}_{\epsilon \sim \mathcal{N}(\epsilon ; 0,1)}\left[\log p_{\theta}\left(x \mid z=\mu_{\phi}(x)+\epsilon * \sigma_{\phi}(x)\right)\right] \\
& \approx \frac{1}{L} \sum_{l=1}^{L} \log p_{\theta}\left(x \mid z^{(l)}=\mu_{\phi}(x)+\epsilon^{(l)} * \sigma_{\phi}(x)\right)
\end{aligned}
$$

The noise is introduced "from outside" the computation graph, and we can evaluate the objective function and take derivatives / get the gradients!

## Reparameterization trick



## ELBO for full data set

You now have all the tools to estimate the ELBO for a whole data set,
$\mathcal{L}(X ; \theta, \phi)=\sum_{n=1}^{N}\left\{\mathbb{E}_{q_{\phi}\left(z^{(n)} \mid x^{(n)}\right)}\left[\log p_{\theta}\left(x^{(n)} \mid z^{(n)}\right)\right]-\operatorname{KL}\left(q_{\phi}\left(z^{(n)} \mid x^{(n)}\right) \| p\left(z^{(n)}\right)\right)\right\}$
take mini-batch subsamples, and use stochastic gradient ascent to maximize it.

## Practical



The end

## Evidence lower bound (ELBQ) for one data point <br> ELBO: can estimate

$$
\begin{aligned}
& \geq \int q_{\phi}(z \mid x) \log \left[\frac{p_{\theta}(x \mid z) p(z)}{q_{\phi}(z \mid x)}\right] \mathrm{d} z \quad[\text { Jensen }] \\
& =\int q_{\phi}(z \mid x) \log p_{\theta}(x \mid z) \mathrm{d} z-\int q_{\phi}(z \mid x) \log \left[\frac{q_{\phi}(z \mid x)}{p(z)}\right] \mathrm{d} z \\
& =\mathbb{E}_{q_{\phi}(z \mid x)}\left[\log p_{\theta}(x \mid z)\right]-\mathrm{KL}\left(q_{\phi}(z \mid x) \| p(z)\right) \\
& \equiv \mathcal{L}(x ; \theta, \phi)
\end{aligned}
$$

